

Bandit Pareto Set Identification in a Multi-Output Linear Model

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Motivation

- Given K **multivariate** distributions (or arms), identify by adaptively sampling them the distributions whose average return is not uniformly worse than any other
- Each arm is associated with some observable features

Applications: clinical trials, large-scale recommender systems, software and hardware design, etc.

Problem setting

- Sub-Gaussian distributions (or arms) over \mathbb{R}^d , ν_1, \dots, ν_K with means (resp.) $\mu_1, \dots, \mu_K \in \mathbb{R}^d$ and **descriptive features** $x_1, \dots, x_K \in \mathbb{R}^h$
- Linearity between vectors means and features i.e. $\mu_i = \Theta^\top x_i$ and $\Theta \in \mathbb{R}^{h \times d}$ is **unknown**
- Large number of arms and a few descriptors: $h \ll K$.

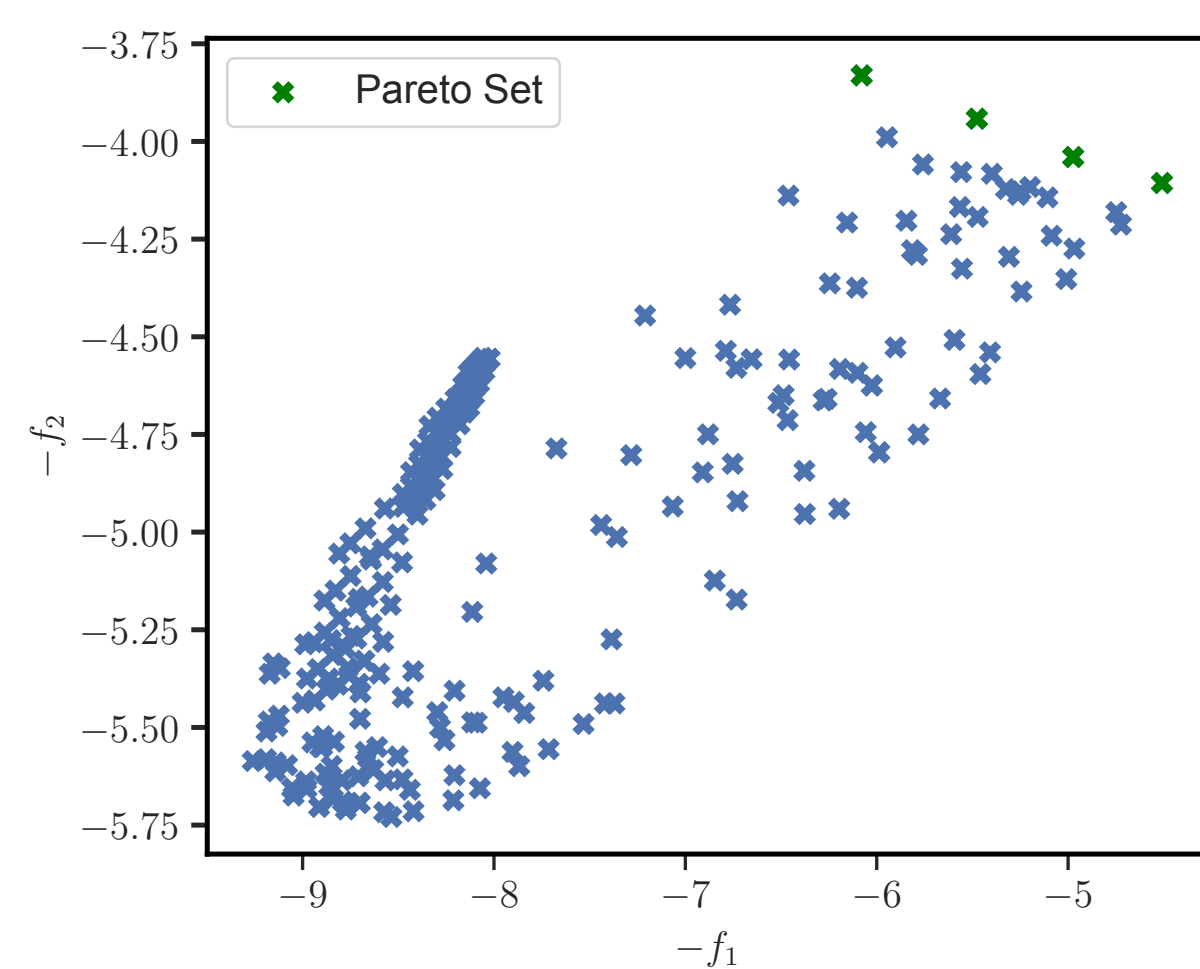


Figure 1: Energy consumption/runtime trade-off of synthesized application-specific networks on chip in Almer et al. [2011]. Each of the 259 architectures is described by 4 features: bus width, buffer depth, CPU clock, internal bandwidth.

Goal: Identify the Pareto Set

$$S^* := \{i \in [K] : \forall j \neq i, \exists c \in \{1, \dots, d\} : \mu_i(c) \geq \mu_j(c)\},$$

with a budget of T samples (**fixed-budget**, e.g., number of chips to synthesize or number of patients in a clinical trial) or with a maximum misidentification rate $\delta \in (0, 1)$ (**fixed-confidence**).

Contribution: we propose an elimination-based optimal-design algorithm for Pareto Set Identification (PSI) in a linear model with nearly-optimal guarantees in fixed-budget and fixed-confidence.

What is the difficulty of classifying an arm ?

For all arms i, j , introduce

$$M(i, j) := \max_{c \in [d]} [\mu_i(c) - \mu_j(c)] \text{ and } m(i, j) := \min_{c \in [d]} [\mu_j(c) - \mu_i(c)],$$

- When $\mu_i \prec \mu_j$, $m(i, j)$ is the minimum increase to a component of μ_i so that it is non-dominated by μ_j ,
- When $\mu_i \not\prec \mu_j$, $M(i, j)$ is the smallest uniform increase of μ_j that makes it dominates μ_i .

For an arm i fixed,

the smaller $M(i, j)$, the “more j is close to be dominating” i (or dominates it by a large margin): i.e for any criterion $c \in [d]$, $(\mu_i(c) - \mu_j(c))$ is small.

the larger $\min_{j \neq i} M(i, j)$, the “more i is optimal”: for any j , there exists a criterion $c_{i,j}$ such that $\mu_i(c_{i,j}) \gg \mu_j(c_{i,j})$

The complexity of unstructured PSI scales as a sum of $1/\Delta_i^2$ terms. For a sub-optimal arm $i \notin S^*$,

$$\Delta_i := \max_{j \in S^*} m(i, j),$$

which is the smallest quantity that should be added component-wise to μ_i to make i appear Pareto optimal w.r.t $\{\mu_i : i \in [K]\}$. For a Pareto-optimal arm $i \in S^*$,

$$\Delta_i := \min(\delta_i^+, \delta_i^-), \text{ where}$$

$$\delta_i^+ := \min_{j \in S^* \setminus \{i\}} [\min(M(i, j), M(j, i))] \text{ and } \delta_i^- := \min_{j \in [K] \setminus S^*} [(M(j, i)_+ + \Delta_j)], \text{ with } (x)_+ := \max(x, 0).$$

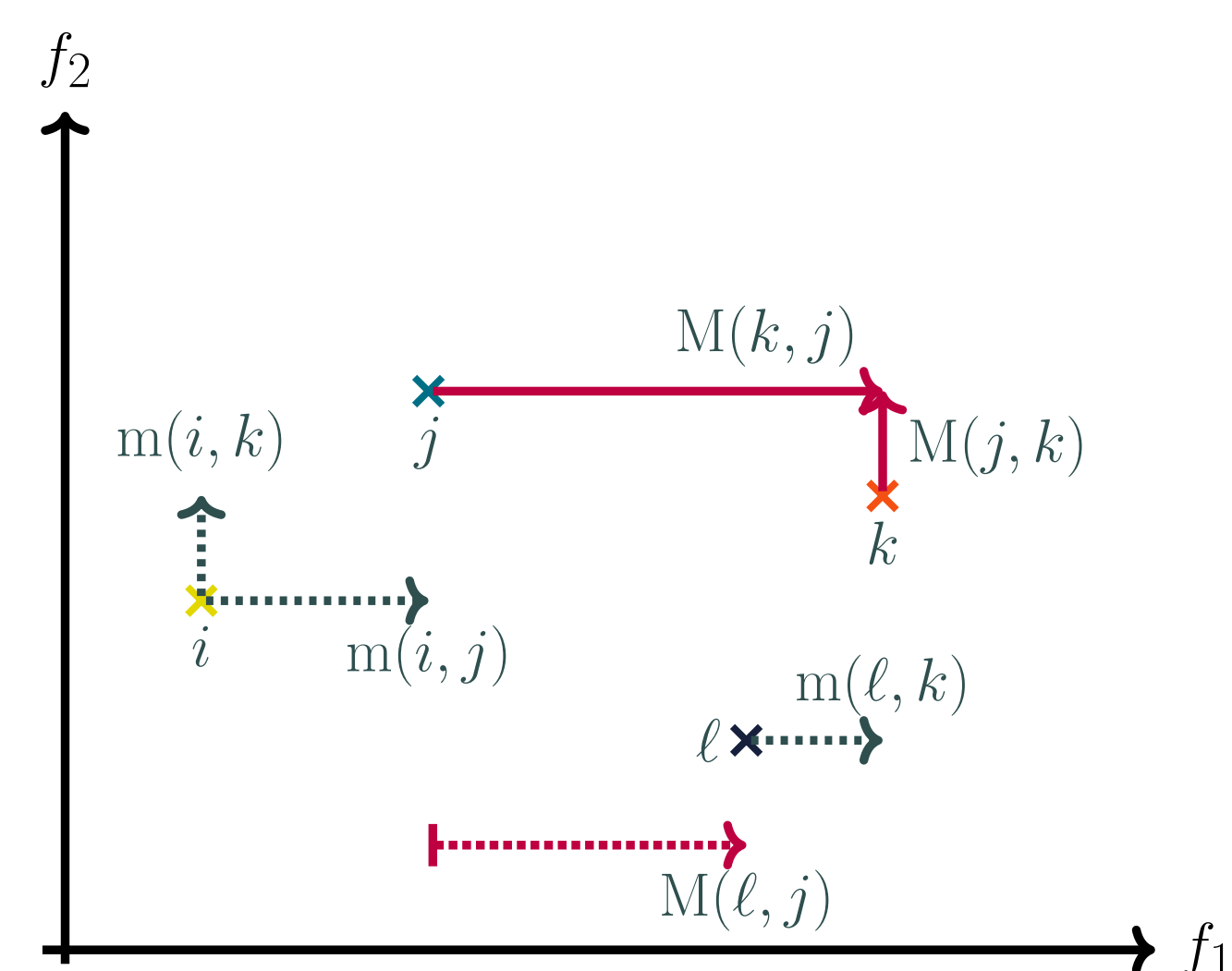


Figure 2: PSI “sub-optimality” gaps. Plain lines represent “distances” between Pareto optimal arms, and dashed lines are for margins from sub-optimal to optimal arms.

- The larger Δ_i the easier it is to classify i as optimal/sub-optimal
- The complexity of **unstructured** PSI is characterized by

$$H_1(\mu) = \sum_{i \in [K]} \frac{1}{\Delta_i^2} \text{ [fixed-confidence] } ; \quad H_2(\mu) := \max_{i \in [K]} \frac{i}{\Delta_i^2} \text{ [fixed-budget] },$$

where $\Delta_1 \leq \Delta_2 \leq \dots \leq \Delta_K$ (w.l.o.g.).

Linear PSI

Lemma 1. By deterministically sampling arms $a_1, a_2, \dots, a_n \in [K]$, and collecting outcomes $y_1 := \Theta^\top x_{a_1} + \eta_1, \dots, y_n := \Theta^\top x_{a_n} + \eta_n \in \mathbb{R}^d$; introducing $X_n := (x_{a_1} \dots x_{a_n})^\top$, $Y_n := (y_1 \dots y_n)^\top$, the estimator $\hat{\Theta}_n$ that minimizes $\mathbb{R}^{h \times d} \ni A \mapsto \|X_n A - Y_n\|_F^2$ satisfies for all $i \in \{a_1, \dots, a_n\}$, $\mathbb{E}[\hat{\Theta}_n^\top x_i] = \mu_i$ and $\text{Cov}(\hat{\Theta}_n^\top x_i) = \|x_i\|_{V_n}^2 \Sigma$. Where $\eta, \eta_1, \dots, \eta_n$ are i.i.d centered, $\text{Cov}(\eta) = \Sigma$ and $V_n := X_n^\top X_n = \sum_{i \in [K]} N_n(i) x_i x_i^\top$; $N_n(i) = \sum_{s \in [n]} \mathbb{I}\{a_s = i\}$.

G-optimal Empirical Gap Elimination [GEGE]

For $r = 1, 2, \dots, \tau$:

- Compute a G-optimal design i.e find

$$\mathbf{w}_r^* \in \arg\min_{\mathbf{w} \in \Delta_{A_r}} \max_{i \in A_r} \|x_i\|_{V_{r, \mathbf{w}}}^2 \text{ with } \Delta_{A_r} : \text{simplex on } A_r \text{ and } V_{r, \mathbf{w}} := \sum_{i \in A_r} \mathbf{w}(i) x_i x_i^\top, \quad (1)$$

- Use (1) and a **rounding** to collect n_r samples, compute $\hat{\Theta}_{n_r}$ and **compute the empirical gaps**

$$\hat{\Delta}_{i, r} := \begin{cases} \hat{\Delta}_{i, r}^* := \max_{j \in A_r \setminus \{i\}} m(i, j; r) & \text{if } i \in A_r \setminus S_r, \\ \min_{j \in A_r \setminus \{i\}} [M(i, j; r) \wedge (M(j, i; r)_+ + (\hat{\Delta}_{i, r}^*)_+)] & \text{else,} \end{cases} \quad (2)$$

where A_r is the set of active arms and S_r , its empirical Pareto Set.

- Discard a **proportion** p_r of the active arms
- Add the discarded empirically Pareto-optimal arms to B_r**

Finally, the algorithm recommends B_τ as the **Pareto Set** and the goal is to upper-bound

$$\mathbb{P}_\mu(B_R \neq S^*) \text{ [fixed-budget] and } \sum_{r=1}^{\tau_\delta} n_r, \text{ with } \mathbb{P}_\mu(B_{\tau_\delta} \neq S^*) \leq \delta \text{ [fixed-confidence]}.$$

Technical Challenge:

- Ensure that the gaps defined in (2) properly estimate the PSI gaps
- Avoid early discarding of some Pareto optimal arms; if not, sub-optimal arms may appear optimal

Theorem 1. In the fixed-budget setting, given a budget $T \geq 45h \log_2 h$, GEGE identifies the Pareto Set with a probability of error smaller than

$$\mathcal{O}\left(\exp\left(-\frac{T}{H_{2, \text{lin}}(\mu) \lceil \log_2(h) \rceil}\right)\right), \text{ with } H_{2, \text{lin}}(\mu) := \max_{i \in [h]} \frac{i}{\Delta_i^2}.$$

In the fixed-confidence setting, given $\delta \in (0, 1)$, with probability $1 - \delta$, GEGE correctly identifies the Pareto Set within $\log_2(2/\Delta_1)$ rounds, using at most

$$\mathcal{O}(H_{1, \text{lin}}(\mu) \log(Kd \log(H_{1, \text{lin}})/\delta) \text{ samples, with } H_{1, \text{lin}}(\mu) = \sum_{i \in [h]} \frac{1}{\Delta_i^2}.$$

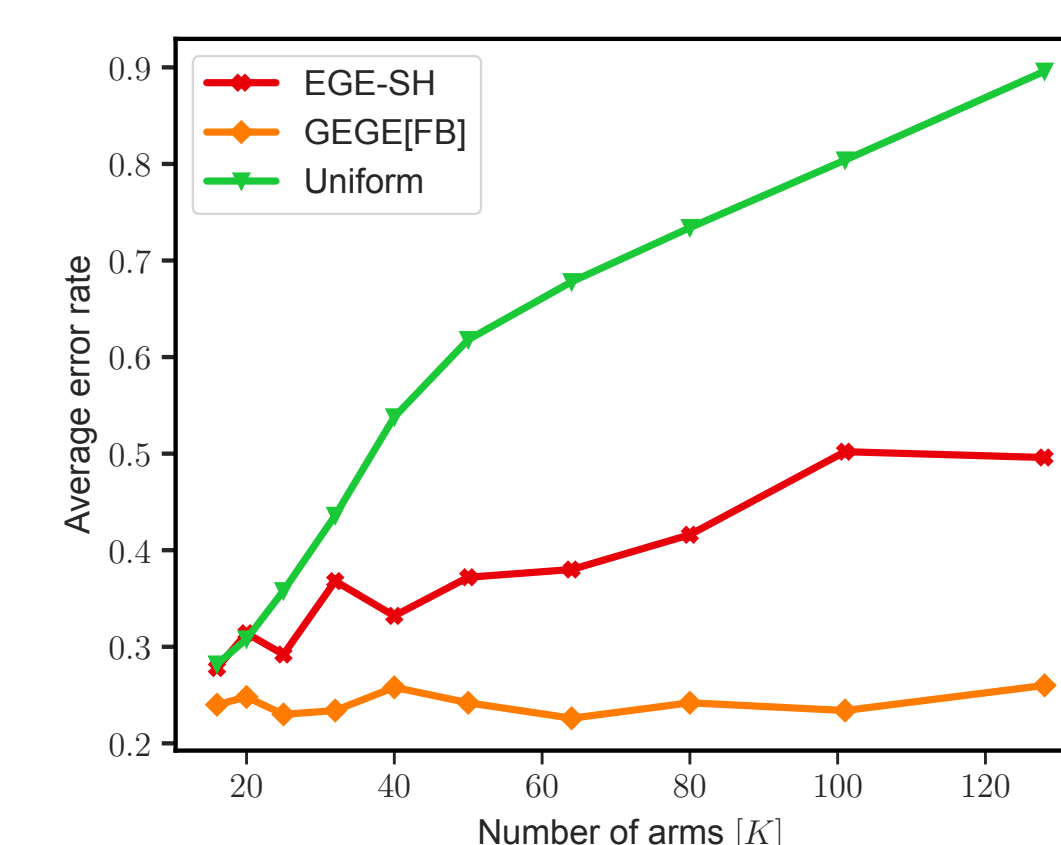
And these guarantees are minimax-optimal.

Some remarks :

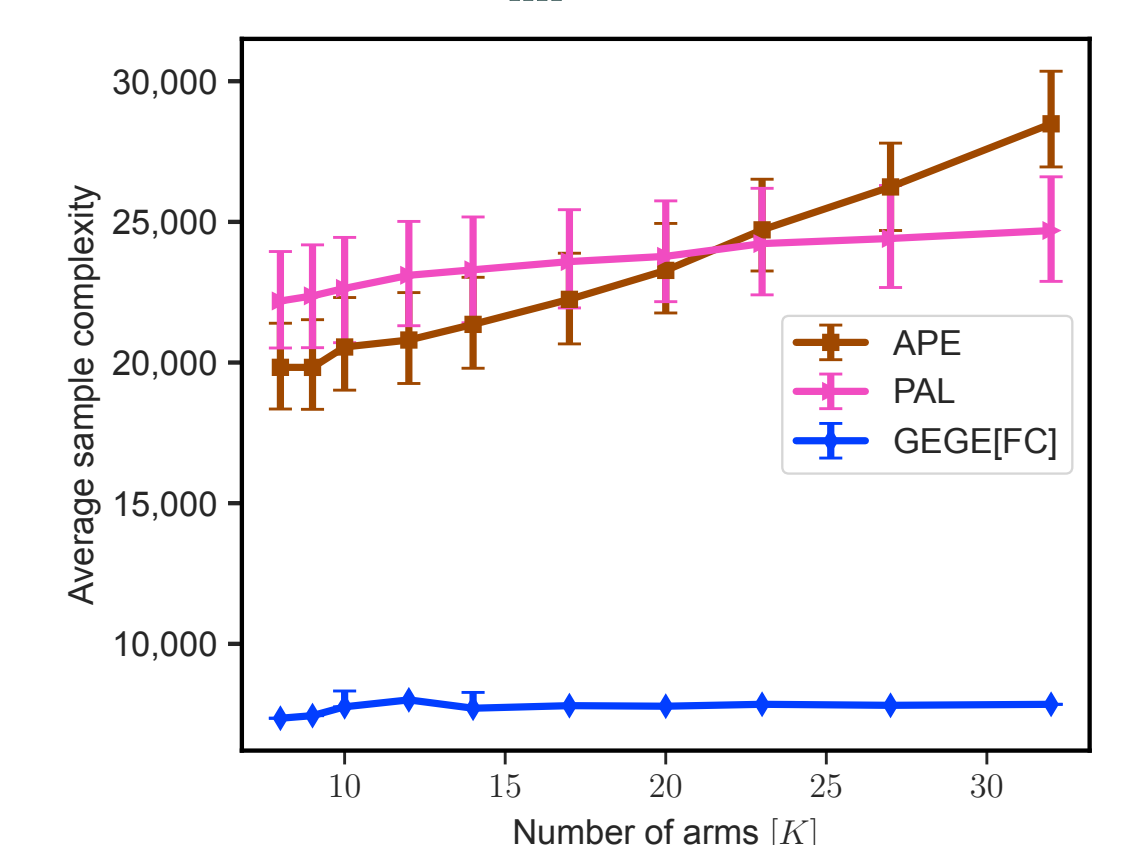
- The complexity of Linear PSI only depends on the gaps of h arms
- $H_{q, \text{lin}}(\mu) \leq H_q(\mu)$ and in some instances, $H_q(\mu)/H_{q, \text{lin}}(\mu) = K/h$ for both $q = 1$ and $q = 2$
- When $h \ll K$, the guarantees of GEGE largely improves upon unstructured algorithms for PSI

Experimental results

- Sample complexity and misidentification rate vs number of arms for H_{lin} 's fixed



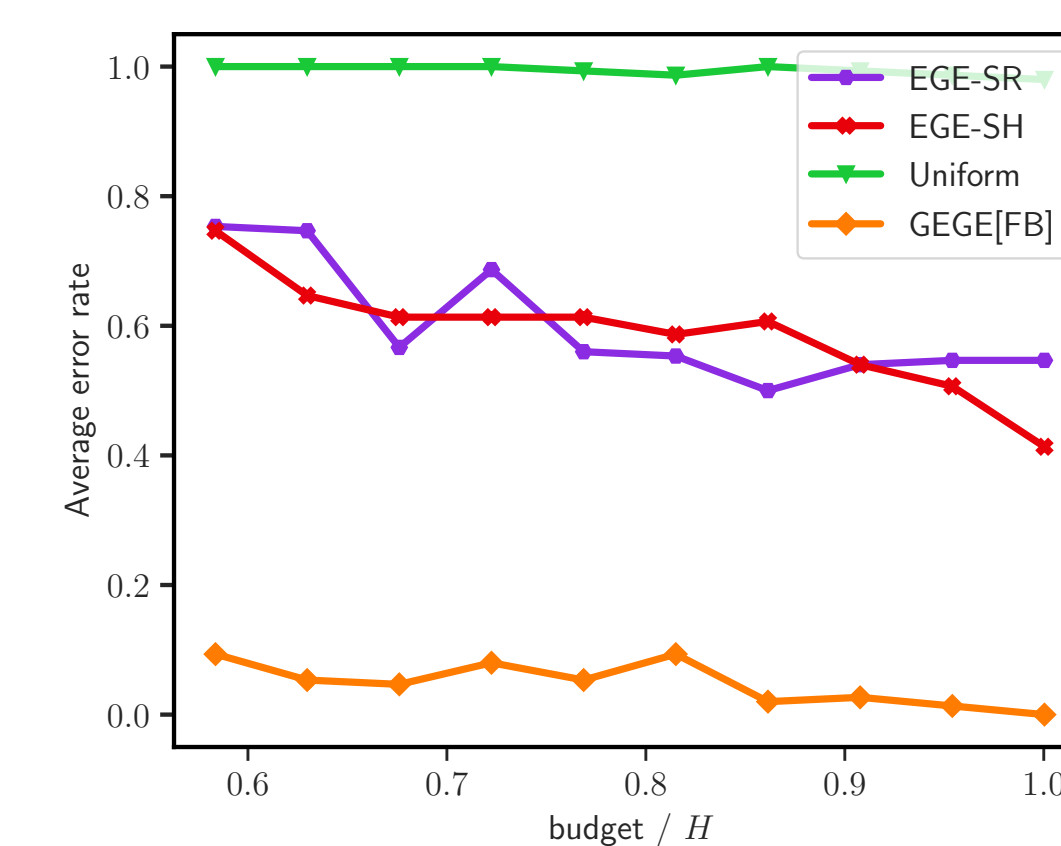
(a) Misidentification rate vs number of arms



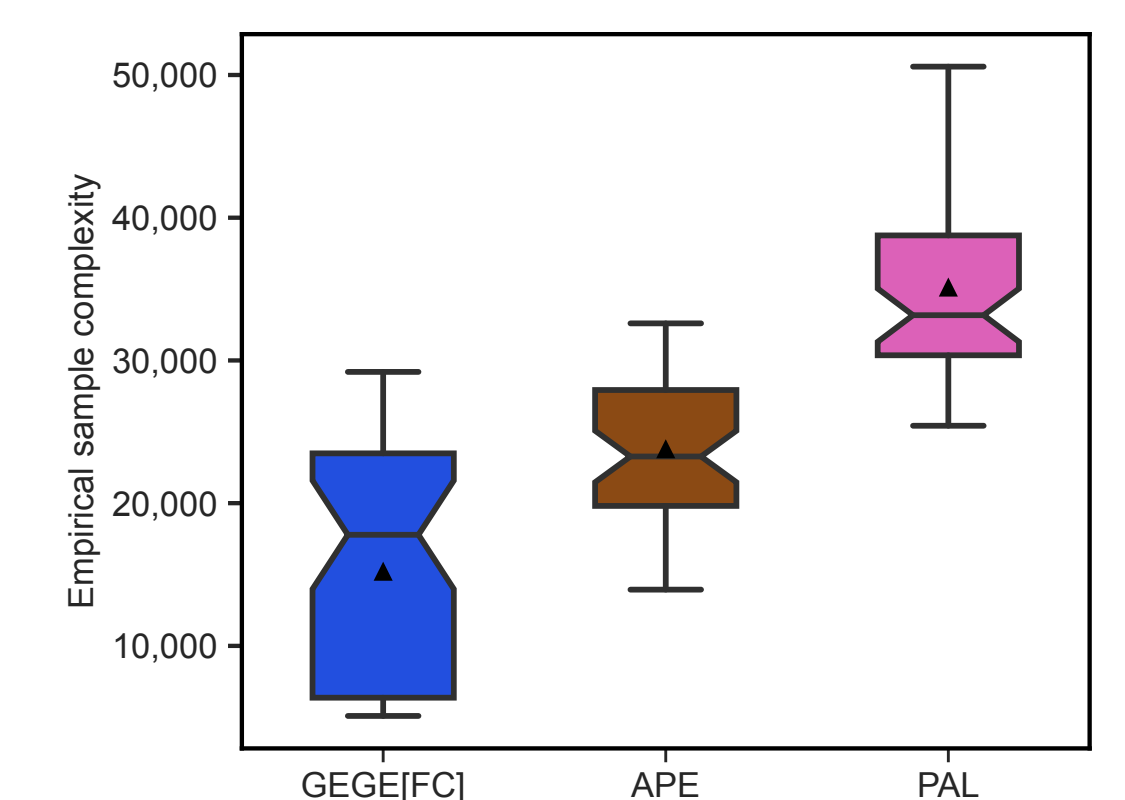
(b) Average sample complexity vs number of arms

Figure 3: Performance vs number of arms. $h = 8, d = 2, \delta = 0.01$ (fixed-confidence). Statistics averaged over 500 runs.

- Performance on the NoC dataset



(a) Misidentification rate vs budget



(b) Sample complexity for $\delta = 0.01$

Figure 4: Benchmark on the NoC dataset averaged over 500 runs. $h = 4, d = 2$ and $\delta = 0.01$ in fixed-confidence.

Conclusion and remarks

- First optimal-design algorithm for PSI in fixed-budget and fixed-confidence regimes
- $H_{1, \text{lin}}(\mu)$ and $H_{2, \text{lin}}(\mu)$ are good proxies to characterize the complexity of linear PSI in the fixed-confidence and fixed-budget settings